



Bayesian Inference 2: Bayesian Model Selection

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MCMC Coffee, Episode 7

Material Reference:

*Statistics, Data Mining, and Machine Learning in
Astronomy*

*Željko Ivezić, Andrew J. Connolly, Jacob T. VanderPlas
& Alexander Gray*

From Last Time

The **Bayes' theorem**

Bayes' rule is not controversial

The frequentist vs. Bayesian controversy sets in when we apply Bayes' rule to the likelihood function $p(D|M)$ to obtain Bayes' theorem:

$$\underbrace{p(M, \theta | D, I)}_{\text{Posterior probability}} = \frac{\underbrace{p(D | M, \theta, I)}_{\text{Likelihood}} \underbrace{p(M, \theta | I)}_{\text{Prior}}}{\underbrace{p(D | I)}_{\text{Evidence}}}$$

D = data

M = model

$\theta = \theta_1, \dots, \theta_k$ parameters of the model

I = any other information

An "**improved belief**" is proportional to the product of an "**initial belief**" and the **probability that the "initial belief" generated the observed data.**



Odds Ratio

- Before we assumed model to be true. What if we can't?
- We consider the odds ratio between two different models we consider

- $$O_{21} \equiv \frac{p(M_2|D, I)}{p(M_1|D, I)}.$$



Evidence and Posterior Probability

- We can integrate over model parameter space spanned by theta

- $$E(M) \equiv p(D|M, I) = \int p(D|M, \theta, I) p(\theta|M, I) d\theta$$

- ^ Evidence or global likelihood
- *Evidence != data. From physics literature*
- With this and prior probability of model, we can get the posterior probability of model M given data D
- These are probabilities, not PDFs!

$$p(M|D, I) = \frac{p(D|M, I) p(M|I)}{p(D|I)},$$



Bayes Factor

- Hardest term to calculate: $p(D|I)$. Drops out though:

$$O_{21} = \frac{E(M_2) p(M_2|I)}{E(M_1) p(M_1|I)} = B_{21} \frac{p(M_2|I)}{p(M_1|I)}.$$

- B_{21} is the Bayes factor

$$B_{21} = \frac{\int p(D|M_2, \theta_2, I) p(\theta_2|M_2, I) d\theta_2}{\int p(D|M_1, \theta_1, I) p(\theta_1|M_1, I) d\theta_1}.$$

- θ_1 and θ_2 can be very different parameter spaces



Interpreting the Odds Ratio

- $O_{21} > 100$: Decisive (100x M2 is more probable than M1)
- $O_{21} > 10$: Strong
- $O_{21} > 3$: Can't say much

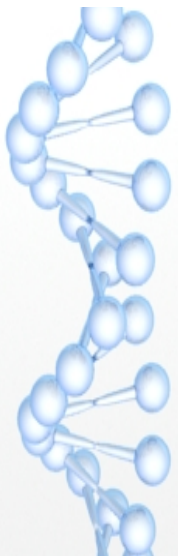
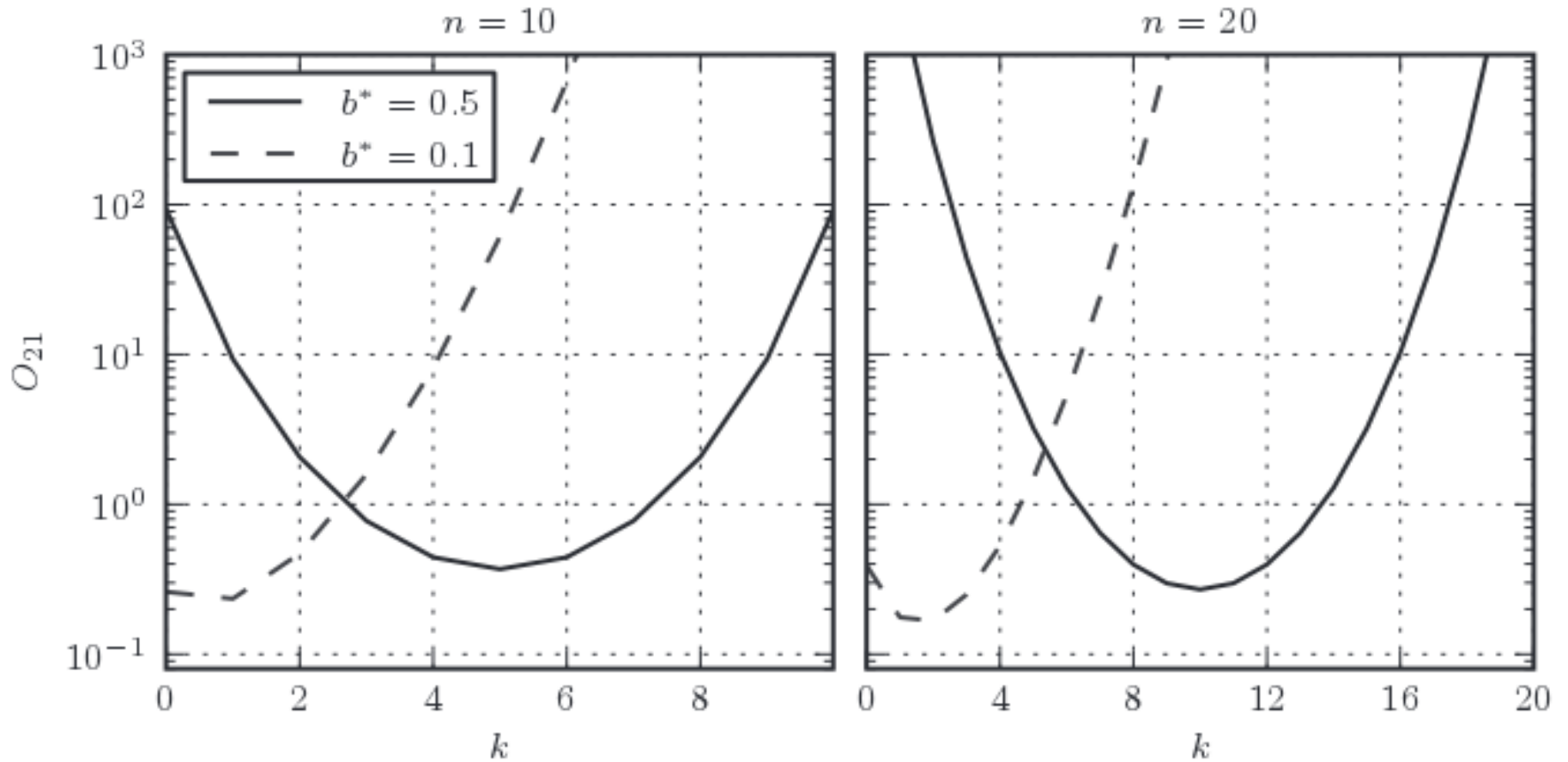
Example: Coin Toss

- M1: known heads probability of b^*
- prior is $\text{del}(b - b^*)$
- Binomial dist:

$$p(k|b, N) = \frac{N!}{k!(N-k)!} b^k (1-b)^{N-k}$$

$$O_{21} = \int_0^1 \left(\frac{b}{b_*}\right)^k \left(\frac{1-b}{1-b_*}\right)^{N-k} db.$$

- M2: b unknown. Uniform prior in 0-1 range
- $O_{21} = \text{sqrt}(\pi/(2N))$, $k = b^*N$, $b=0.5$
- $N=10000$ needed for 1% precision with odds 1/80 for O_{21}



$b^* = 0.5$ (fair) and $b^* = 0.1$.

If $k/N = 0.1$, O_{21} is ~ 9 for $N=10$,

$O_{21} \sim 263$ for $N=20$

If $b^* = k/N$ and $b^* = 0.5$ we need $N > 157$ for strong evidence

$N=10000$ needed for 1% precision with odds 1/80



Hypothesis Testing

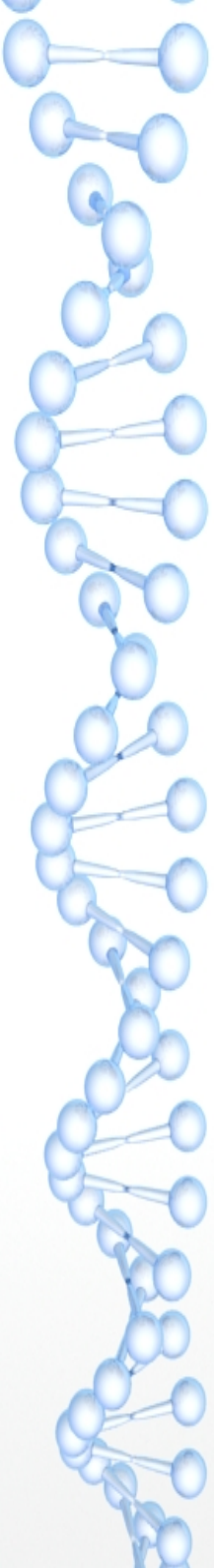
- In classical, we reject a null hypothesis based on confidence intervals
- In Bayesian, we cannot reject without alternate hypothesis!
- Let's consider $p(M_1) + p(M_2) = 1$. M_1 is “null” hypothesis
- Equal priors:
- $$O_{21} = B_{21} = \frac{p(D|M_1)}{p(D|M_2)}$$
- For coin, let's say that $k=16$ heads out of $N=20$ tries
- If $\sigma=2.24$, it is $2.7 \cdot \sigma$ away, above 0.05 confidence interval leading us to reject.
- In Bayesian, we have alternate hypothesis that there is an unknown probability
- By comparing, we can O_{21} of 10 in favor of an unfair coin

Occam's Razor

- Choose the simplest model in fair agreement with data
- Let's say prior pdf is flat in range:
 - $-\Delta_\mu/2 < \mu < \Delta_\mu/2$, and thus $p(\mu|I) = 1/\Delta_\mu$.
- Data is more informative than prior when $\sigma_\mu \ll \Delta_\mu$ for Gaussian data with width σ_μ
- Global likelihood:
 - $$E(M) \approx \sqrt{2\pi} L^0(M) \frac{\sigma_\mu}{\Delta_\mu}.$$
- $E \ll L^0$ when $\sigma_\mu \ll \Delta_\mu$

Occam's Razor

- Each parameter has a multiplicative penalty proportional to σ/Δ , where $\sigma \sim \Delta$ gives no penalty (unconstrained)
- The odds ratio can justify an additional parameter if the penalty is offset a larger maximum likelihood L_0 or by prior model probability ratio $p(M_2|I) / p(M_1|I)$





Occam's Razor

- For coin flip example from what we had before:

- $$O_{21} = \frac{E(M_2)}{E(M_1)} \approx \sqrt{2\pi} \sigma_b \left(\frac{b^0}{b_*} \right)^k \left(\frac{1 - b^0}{1 - b_*} \right)^{N-k},$$

As M_2 has a free parameter, we can see its favor decrease with posterior PDF width (sigma)

- It increases if b_0 is far from b^*